Sampling in Python by datacamp

population v sample population is the complete dataset *tricksy thing - most of the time in statistics we won't know what the whole population is like sample is the subset of data that we are working with

Example - let's consider the relationship between cup points and flavor in an evaluation of coffee pull our entire population whcih in this case is 1,338 coffee ratings pts_vs_flavor_pop = coffee_ratings[['total_cup_point', 'flavor']] lets pull a sample pts_vs_flavor_samp = pts_vs_flavor_pop.sample(n=10) set n to our desired sample size sample method in pandas package sample method returns a random subset of rows **by default rows cannot appear multiple times, so we are guaranteed to have ten unique rows in our sample

Can also use the sample method on pandas Series by using subsetting cup_poins_samp = coffee_ratings['total_cup_points'].sample(n=10)

Population parameter is a calculation made on the population dataset import numpy as np np.mean(pts_vs_flavor_pop['total_cup_points']) with pandas pts_vs_flavor_pop['flavor'].mean()

For a sample these parameters are called: sample statistic or point estimate np.mean(cup_points_samp) pts_vs_flavor_samp['flavor'].mean()

Convenience sampling collecting data by the easiest method often prone to sample bias sample bias sample is not representative of population **randomness is one way to help avoid this

Visualizing selection bias coffee_ratings['total_cup_points'].hist(bins=np.arange(59, 93, 2)) plt.show() #we looked at the dataset and subjectively chose a low rating of 59 and a high of 91 #the 2 creates bins of width 2

Pseudo-random number generation

randomness really mean

If we want to choose data points at random from a population, we shoudn't be able to predict which data points would be selected ahead of time in some systematic way.

we use pseudo-random number generation because it is cheap and fast next 'random' number is actually calculated from previous 'random' number the first 'random' number is calculated from a 'seed' value **same seed value yields the same random numbers

Random number generating functions

• Prepend with numpy.random, such as numpy.random.beta()

Visualizing random numbers

randoms = np.random.beta(a=2, b=2, size=5000)

plt.hist(randoms, bins=np.arange(0,1, 0.05)

plt.show()

#a,b arguments to the beta function specify distribution parameters

Setting a random seed np.random.seed(any integer)

another example of random number generator function np.random.normal(loc=2, scale=1.5, size=2) this one generates pseudo-random numbers from the normal distribution loc and scale arguments set the mean and standard deviation of the distribution

Simple random and systematic sampling simple random sampling is like a raffle or lottery also called SRS sometimes coffee_ratings.sample(n=t, random_state=190000113) random_state allows us to set a seed within the sample request

```
systematic sampling
samples the population at regular intervals
example would be taking every fifth sample
harder but still possible to do with pandas
example - sampling 5 coffees from our coffee set
sample_size = 5pop_size = len(coffee_ratings)
print(pop_size)
output > 1338
interval = pop_size // sample_size
#// represents integer division, like standard division but discards any fractional 
part
interval = 1338 // 5 = 267 instead of 267.6
to select every 267th coffee
coffee_ratings.iloc[::interval]
pass :: to tell panda to select every 267th coffee till the end of the DF
```
The trouble with systematic sampling can accidentally introduce bias into the statistics that we calculate important to visualize data prior to attempt to avoid these types of bias example in the coffee set coffee_ratings_with_id = coffee_ratings.reset_index() coffee_ratings_with_id.plot(x='index', y='aftertaste', kind='scatter') plt.show()

The plot clearly shows a bias as the samples go from 0 to 1300 *systematic sampling is only safe if we don't see a pattern in this scatter plot

Making systematic sampling safe

we can randomize the row order before sampling

shuffled = coffee_ratings.sample(frac=1)

frac argument lets us specify the proportion of the dataset to return in the sample rather than the absolute number of rows than n specifies

**setting frac=1, randomly samples the whole dataset

Next, we need the indices to be reset so that they go in order from zero again dropt argument set to True clears the previous row indexes

then chaining them to another reset_index call creates a column containing these new indexes

shuffled = shuffled.reset_index(drop =True).reset_index() shuffled.plot(x='index', y='aftertaste', kind='scatter') plt.show()

***biggest key point is this whole process is essentially the same as simple

random sampling

Stratified and weighted random sampling

stratified sampling

a technique that allow us to sample a population that contains subgroups example - working with coffees via their country

to see how the data is spersed (often useful to know where the most data is)

top_counts = coffee_ratings['country_of_origin'].value_counts()

top_counts.head(6)

then filter for the 6 top countries

top_counted_countries = ['Mexico', 'Columbia', 'Guatemala', 'Brazil', 'Taiwan',

'United States (Hwaii)']

top_counted_subset =

coffee_ratings['country_of_origin'].isin(top_counted_countries)

coffee ratings top = coffee ratings [top_counted_subset]

now take a percent of the simple random sample

do this using the frac argument 0.1 = 10%

coffee_ratings_samp = coffee_ratings_top.sample(frac=0.1, random_state=2021) look at the counts now

coffee_ratings_samp['country_of_origin'].value_counts(normalize=True) set normalize argument to True to convert the counts into a proportion this proportion shows what proportion of coffees in the sample came from each country

**if we care about the proportions of each country in the sample closely matching those in the population, then we can group the data by country before taking the simple random sample

```
coffee_ratings_strat = coffee_ratings_top.groupby('country_of_origin') \
     .sample(frac=0.1, random_state=2021)
```
*just practicing with the '\', it is there simple to sparse out code to make it more readable

**now the proportions of each country in the stratified sample are much closer to those in the population

*a variation of stratified sampling is to sample equal counts from each group to do this:

```
coffee_ratings_eq = coffee_ratings_top.groupby('country_of_origin').sample(n=15, 
random_state=2021)
```
the difference is using 'n' instead of 'frac'

Weighted random sampling

specifies weights to adjust the relative probability of a row being sampled example

say we wanted to have a higher proportion of Taiwanese coffees in the sample than in the population

import numpy as np coffee_ratings_weight = coffee_ratings_top condtion = coffee_ratings_weight['country_of_origin'] == 'Taiwan' coffee_ratings_weight['weight'] = np.where(condition, 2, 1) **where function - here we have set a weight of 2 to 1 when 'condition' is met **here this will weight Taiwanese coffee 2 to 1 ie Taiwanese coffee has two times the chance of being picked compared to the other coffees now we have to use the 'weights' argument in our 'sample' method coffee_ratings_weight = coffee_ratings_weight.sample(frac=0.1, weights='weight') **this sort of weighted sampling is common in political polling, where we need to correct for under or over representation of demographic groups

Example

Proportion of employees by Education level education_counts_pop = attrition_pop['Education'].value_counts(normalize=True)

Print education_counts_pop print(education_counts_pop)

```
# Proportional stratified sampling for 40% of each Education group
attrition_strat = attrition_pop.groupby('Education')\
   .sample(frac=0.4, random_state=2022)
```
Calculate the Education level proportions from attrition_strat education_counts_strat = attrition_strat['Education'].value_counts(normalize=True)

```
# Print education counts strat
print(education_counts_strat)
```
Weighted example # Plot YearsAtCompany from attrition_pop as a histogram attrition_pop['YearsAtCompany'].hist(bins=np.arange(0, 41, 1)) plt.show()

```
# Sample 400 employees weighted by YearsAtCompany
attrition_weight = attrition_pop.sample(n=400, weights="YearsAtCompany")
```

```
# Plot YearsAtCompany from attrition_weight as a histogram
attrition_weight['YearsAtCompany'].hist(bins=(np.arange(0, 41, 1)))
plt.show()
```
Cluster sampling

-use SRS to pick some subgroups -use SRS on only those subgroups in contrast stratified sampling -splits the population into subgroups -then uses SRS on every subgroup clustering is often used to cut down on costs example - coffee dataset first stage of cluster sampling is to randomly cut down the number of varieties (or unique values) import random #create a list varieties_pop = list(coffee_ratings['variety'].unique()) varieties_samp = random.sample(varieties_pop, k=3) we specify how many unique varieties that we want in this example with the 'k' argument stage 2 - perform SRS on each of the three varieties #filter dataset for rows where the variety is one of the three selected using the .isin() method variety_condition = coffee_ratings['variety'].isin(varieties_samp) coffee_ratings_cluster = coffee_ratings[variety_condition] #then ensure the isin() method removes levels with zero rows, we apply the .cat.remove_unused_categories() on the focal Series **this is important or otherwise an error may occur when sampling by variety level coffee_ratings_cluster['variety'] = coffee_ratings_cluster['variety'].cat.remove_unused_categories() then groupby and sample as we've done in previous examples coffee_ratings_cluster.groupby('variety').sample(n=5, random_state=2021) **we randomly sampled the subgroups to include, then we randomly sampled rows from those subgroups

Example

Create a list of unique JobRole values job_roles_pop = list(attrition_pop['JobRole'].unique())

Randomly sample four JobRole values $job_roles_ samp = random.sample(job_roles_pop, k=4)$

Filter for rows where JobRole is in job_roles_samp jobrole_condition = attrition_pop['JobRole'].isin(job_roles_samp) ***we then apply the 'jobrole_condition' as a filter on the 'attrition_pop' DF using boolen indexing ***this filters out the rows where the 'JobRole' is not in the sampled list attrition filtered = attrition pop[jobrole condition]

```
# Remove categories with no rows
attrition_filtered['JobRole'] = 
attrition_filtered['JobRole'].cat.remove_unused_categories()
```
Randomly sample 10 employees from each sampled job role attrition_clust = attrition_filtered.groupby('JobRole').sample(n=10, random_state=2022)

Print the sample print(attrition_clust)

Another example # Create a list of unique RelationshipSatisfaction values satisfaction_unique = list(attrition_pop['RelationshipSatisfaction'].unique())

Randomly sample 2 unique satisfaction values satisfaction_samp = random.sample(satisfaction_unique, k=2)

Filter for satisfaction_samp and clear unused categories from RelationshipSatisfaction satis_condition = attrition_pop['RelationshipSatisfaction'].isin(satisfaction_samp) attrition_clust_prep = attrition_pop[satis_condition] attrition_clust_prep['RelationshipSatisfaction'] = attrition_clust_prep['RelationshipSatisfaction'].cat.remove_unused_categories()

Perform cluster sampling on the selected group, getting 0.25 of attrition_pop attrition $clust =$ attrition clust prep.groupby('RelationshipSatisfaction').sample(n=367, random_state=2022)

Relative error of point estimates

ie how the size of the sampe affects the accuracy of the point estimates len function pulls the number of rows in the sample ie the number of observations in the sample **larger sample sizes usually give us more accurate results

Relative error is the absolute (ignore minus signs) difference between the population and a sample mean population parameter: population_mean = coffee_ratings['total_cup_points'].mean() point estimate:

sample_mean = coffee_ratings.sample(n=sample_size)['total_cup_points'].mean() relative error as a percentage:

rel_error_pct = 100 * abs(population_mean - sample_mean) / population_mean

Plotting relative error vs sample size import matplotlib.pylplot as plt errors.plot(x='sample_size', y='relative_error', kind='line') plt.show()

really noisy, particularly for small samples amplitude and noise decrease with more samples **relative error decreases to zero when the sample size = population

example

Generate a simple random sample of 100 rows, with seed 2022 attrition_srs100 = attrition_pop.sample(n=100, random_state=2022)

Calculate the mean employee attrition in the sample mean_attrition_srs100 = attrition_srs100['Attrition'].mean()

Calculate the relative error percentage rel_error_pct100 = 100 * abs(mean_attrition_pop - mean_attrition_srs100) / mean_attrition_pop

Print rel_error_pct100 print(rel_error_pct100)

Creating a sampling distribution

using a for loop to run the same code many times **this example samples 30 random coffees from the coffee_ratings dataset, calculates the mean, then puts the result (called a replicate) into the open list, does this a 1000 times mean_cup_points_1000 = $\lceil \cdot \rceil$ for i in range(1000) mean_cup_points_1000.append(coffee_ratings.sample(n=30) ['total_cup_points'].mean())

print(mean_cup_points_1000)

output > a thousand separate sample means

Visualize the distribution **again, often the best plot to visualize a distribution is a histogram plt.hist(mean_cup_points_1000, bins=30) plt.show() this example creates a nice bell shape curve (normal distribution) **a distribution of replicates of point estimates is called a 'sampling distribution' decreasing the sample size increases the range of the x-values on the histogram inverse increasing the sample size decreases the range of the x-values

*increasing the number of replicates didn't affect the relative error of the sample means but it did resultin a more consistent shape to the distribution

Approximate sampling distributions **we can generate all possible combinations of rolls using the pandas expand_grid function uses the itertools package example with four dice dice = expand_grid({'die1': [1,2,3,4,5,6], 'die2': [1,2,3,4,5,6], 'die3': [1,2,3,4,5,6], 'die4': [1,2,3,4,5,6]}) this is 6^4 or 1296 dice roll combinations now consider the 'mean roll' add a column to the DataFrame dice['mean_roll'] = $(dice['die1'] + dice['die2'] + dice['die3'] + dice['die4']) / 4$ **since the mean roll takes discrete values instead of continuous values, the best way to see the distribution of mean roll is to draw a bar plot **we are interested in the counts of each value so this is better suited as a data type category

dice['mean_roll'] = dice['mean_roll'].astype('category')

dice['mean_roll'].value_counts(sort=False).plot(kind='bar')

sort argument to False ensures the x-axis ranges from one to six instead of sorting the bars by frequency

*interesting fact, with just a 100 die we can get a number of possible outcomes equivalent to the number of atoms in the universe

this obviously shows how quickly we can enter the computationally impossible **what this means is that we have to rely on approximations

Simulating - example simulating the mean of four dice rolls

```
sample_means_1000 = []
```
for i in range(1000):

sample_means_1000. append(

```
np.random.choice(list(range(1,7)), size=4, replace=True).mean())
```
the choice method allow you to randomly select elements from an array based on specified probabilities

first argument is the sequence or array from where the randome samples are to be drawn from

size argument specifies the shape of the output (in our example we are using 4 die)

replace argument to True states each element can be sampled more than once *can add 'p' argument at end to give individualized weights/probabilities this simulation gives us an approximate sampling distribution

Example variant and all together

Expand a grid representing 5 8-sided dice

```
dice = expand_grid(
```
 {'die1': [1, 2, 3, 4, 5, 6, 7, 8], 'die2': [1, 2, 3, 4, 5, 6, 7, 8], 'die3': [1, 2, 3, 4, 5, 6, 7, 8], 'die4': [1, 2, 3, 4, 5, 6, 7, 8], 'die5': [1, 2, 3, 4, 5, 6, 7, 8] })

```
# Add a column of mean rolls and convert to a categorical
dice['mean_roll'] = (dice['die1'] + dice['die2'] + dice['die3'] + dice['die4'] + 
              dice['die5']) / 5
dice['mean_roll'] = dice['mean_roll'].astype('category')
```

```
# Draw a bar plot of mean roll
dice['mean_roll'].value_counts(sort=False).plot(kind='bar')
plt.show()
```
Central Limit Theorem

**means of independent samples have normal distributions as the sample size increases we see the distribution becoming more Gaussian (ie Normal) and the width of the sampling distribution gets narrower

Population and sampling distribution means and standard deviations in our coffee_ratings example we took four separate sample means (roughly 5, 20, 80, and 320)

creating four separate distributions

however all four sample means are relatively close to the population mean

population std for this example is ~2.7

sample std's are relatively smaller (roughly 1.2, .6, .3, .13)

**std decreases as the sample size increases

why is this?

***first remember when calculating std with pandas, you have to set ddof argument to 0 when calculating population std

and ddof to 1 (sample std is pandas default) when calculating std for a sample another consequence of the CLT is that if we divide the population std (in this case 2.7) by the square root of the sample size we will get an estimate of the std of the sampling distribution for that sample size

this isn't exact because of the randomness involved in the sampling process, but its pretty close

***standard error is the std of the sampling distribution

useful in estimating population std to setting expectations on what level of variability we would expect from the sampling process

Introduction to bootstrapping

example - focus on sampling with resampling

coffee_focus = coffee_ratings[['variety', 'country_of_origin', 'flavor']]

#to see which rows ended up in the sample, add a row index column called index with the reset_index method

coffee_focus = coffee_focus.reset_index()

#call sample() but this time with argument replace=True

coffee_resamp = coffee_focus.sample(frac=1, replace=True)

#count the values of the index column to see how many times each coffee ended up in the resampled dataset

coffee_resamp['index'].value_counts()

**this means that some coffees were used multiple times while others were not used at all

num_unique_coffees = len(coffee_resamp.drop_duplicates(subset='index')) #find out how many coffees weren't sampled at all

len(coffee ratings) - num unique coffees

Bootstrapping

we use resampling for a technique called bootstrapping

bootstrapping can be seen in some ways as the opposite of sampling from a population

with regular sampling we treat the dataset as the population and move to a smaller sample

**with bootstrapping we treat the dataset as a sample and use it to build up a theoretical population

when to use bootstrapping:

to try to understand the variability due to sampling

theoretically allows us to develop an understanding of sampling variability using a single sample

important in cases where we aren't able to sample the population multiple times to create a sampling distribution

Bootstrapping process

1. make a resample of the same size as the original sample

- 2. calculate the statistic of interest for this bootstrap sample
- 3. repeat steps 1 and 2 many times

**the resulting statistics are bootstrap statistics, and they form a bootstrap distribution

```
For loop this bootstrapping son of a gun
```

```
mean_flavors_1000 = []
```
for i in range(1000):

```
mean_flavors_1000.append(np.mean(coffee_sample(frac=1, replace=True)
['flavor']))
```

```
plt.hist(mean_flavors_1000,)
```

```
plt.show()
```

```
example
# Replicate this 1000 times
mean_danceability_1000 = []
for i in range(1000):
   mean_danceability_1000.append(
      np.mean(spotify_sample.sample(frac=1, replace=True)['danceability'])
  \left( \right)
```
Draw a histogram of the resample means plt.hist(mean_danceability_1000) plt.show()

Comparing sampling and bootstrap distributions

remember in the bootstrap distribution that each value is an estimate of the mean flavor score

recall that each of these values corresoponds to one potential sample mean from the thoeretical population

taking the mean of those means gives us a guess at the population mean this mean will often be very close to the true population mean, but there is a key difference:

*** key key element the bootstrop mean is usually almost identical to its original sample mean

**this can be a bad thing if the original sample wasn't closely representative of the population

this means that our bootstrap distribution mean will not be a good estimate of our true population mean

**bootstrapping cannot correct potential biases due to differences between the sample and the population

calculating the std of our bootstrapping distribution vs our sample we get an entirely different number

why?

remember that one goal of bootstrapping is to quantify what variability we might expect in our sample statistic as we go from one sample to another

recall that this quantity is called the standard error as measured by the std of the sampling distribution of that statistic

std of the bootstrap means can be used as a way to estimate this measure of uncertainty

if we multiply that SE by the sqrt of the sample size, we get an estimate of the std in the original population

standard $error = np.stdoutbootstrap\$ distn, ddof=1)

standard error is the std of the statistic of interest

standard_error * np.sqrt(samplesize)

standard error times square root of sample size estimates the population standard deviation

estimated standard error > std of the bootstrap distribution for a sample statistic again pop std approx= SE x sqrt(samplesize)

***although bootstrapping was poor at estimating the population mean, it is generally great for estimating the population std

```
example
# Calculate the population std dev popularity
pop_sd = spotify_population['popularity'].std(ddof=0)
```

```
# Calculate the original sample std dev popularity
samp_sd = spotify_sample['popularity'].std()
```

```
# Calculate the sampling dist'n estimate of std dev popularity
samp distn sd = np.std(sampling distribution, ddof=1) * (np.sqrt(5000))
```

```
# Calculate the bootstrap dist'n estimate of std dev popularity
boot_distn_sd = np.std(bootstrap_distribution, ddof=1) * (np.sqrt(5000))
```

```
# Print the standard deviations
print([pop_sd, samp_sd, samp_distn_sd, boot_distn_sd])
```
Confidence intervals values within on standard deviation of the mean this gives a good sense of where most of the values in a distribution lie rough example - weather prediction we believe tomorrow will be approx 47F we report a range of 40-54, this is a confidence interval written as 47F (40F,54F) or 47F [40F, 54F] or 47 +-7F 47F is the 'point estimate' with our estimate we have determined a margin of error of 7F

```
example 
get mean then +- std
np.mean(coffee_boot_distn) - np.std(coffee_boot_distn, ddof=1)
np.mean(coffee_boot_distn) + np.std(coffee_boot_distn, ddof=1)
```

```
Quantile method for confidence intervals
quantiles split distributions into sections containing a proportion of the total data
we want 95% of values
go from 0.025 quantile to 0.975 quantile
to calculate the lower and upper bounds for this CI:
np.quantile(coffee_boot_distn, 0.025)
np.quantile(coffee_boot_distn, 0.975)
```
Second method to calculate confidence intervals normal distribution's inverse cummulative distribution function normal distribution's bell curve is the probability density function (PDF) using calculus we can integrate this to get area under bell curve this forms the cumulative distribution function (CDF) by flipping the x and y axes, we get the inverse CDF

from scipy.stats import norm

norm.ppf(quantile, loc=0, scale=1)

takes a quantile between 0 and 1 and returns the values of the normal distribution for that quantile

default parameter of loc is 0 and scale is 1 which corresponds to the standard normal distribution

**notice in the above graph, the values corresponding to 0.025 and 0.975 are about minus and plus 2 for the standard normal distribution

this second method is called the standard error method for CI #first calculate the point estimate, which is the mean of the bootstrap distribution point_estimate = np.mean(coffee_boot_distn)

#second calculate the standard error, which is estimated by the std of the bootstrap distribution

```
std_error = np.std(coffee_boot_distn, ddof=1)
```
#then we call .ppf to get the inverse CDF of the normal distribution with the same mean and std as the bootstrap distribution

from scipy.stats import norm

```
lower = norm.ppf(0.025, loc=point_estimate, scale=std_error)
```

```
upper = norm.ppf(0.975, loc=point_estimate, scale=std_error)
```

```
print((lower, upper))
```
Per lecturer, the most important things

-std of a bootstrap distribution statistic is a good approximation of the standard error of the sampling distribution

-the normal distribution tends to be a good approximation for bootstrap distributions